

# Linear models for regression

Lecture 2

Simplest model:

$$y(\vec{x}, \vec{w}) = w_0 + \sum_{j=1}^D w_j x_j$$

$\{\vec{x}_1, \dots, \vec{x}_N\}$  observations  $\rightarrow$  each  $\vec{x} = \{x_1, \dots, x_D\}$   
 $\{t_1, \dots, t_N\}$  target variables

More generally,

$$y(\vec{x}, \vec{w}) = w_0 + \sum_{j=1}^D w_j g_j(\vec{x})$$

bias       $\nwarrow$   
 prm       $\swarrow$

still a linear model  
 because  $y$  is linear in  $\vec{w}$

Sometimes, one defines  $g_0(\vec{x}) = 1$ , s.t.

$$y(\vec{x}, \vec{w}) = \sum_{j=0}^D w_j g_j(\vec{x}) = \vec{w}^\top \vec{g}(\vec{x})$$

$$\vec{w} = (w_0, \dots, w_D) \quad \vec{g} = (g_0, \dots, g_D)$$

Popular choices of basis f's:

(1D examples)

(1) Gaussian  $g_j(x) = e^{-\frac{(x-\mu_j)^2}{2s^2}}$

(2) Sigmoidal  $g_j(x) = \sigma\left(\frac{x-\mu_j}{s}\right)$ , where

$$\sigma(a) = \frac{1}{1+e^{-a}}$$

As before, we assume that

$$t = y(\vec{x}, \vec{w}) + \xi$$

↑ target      ↑ model      ↑ noise

$$p(\xi | \beta) = \mathcal{N}(0, \beta^{-1})$$

Correspondingly,

$$p(t | \vec{x}, \vec{w}, \beta) = \mathcal{N}(t | y(\vec{x}, \vec{w}), \beta^{-1})$$

Now consider  $\vec{X} = \{\vec{x}_1, \dots, \vec{x}_N\}$   
 $\vec{t} = \{t_1, \dots, t_N\}$   
target vars

Under the independence assumption,

$$P(\vec{t} | \vec{X}, \vec{w}, \beta) = \prod_{n=1}^N \mathcal{N}(t_n | \vec{w}^T \vec{g}(\vec{x}_n), \beta^{-1})$$

drop  
 conditional dependence  
 on  $\vec{X}$  for brevity

Then

$$\log P(\vec{t} | \vec{w}, \beta) = \frac{N}{2} \log \beta - \frac{N}{2} \log(2\pi) - \beta \underbrace{\frac{1}{2} \sum_{n=1}^N (t_n - \vec{w}^T \vec{g}(\vec{x}_n))^2}_{E_\beta(\vec{w})}$$

$\sum_{n=1}^N \log \mathcal{N}(t_n | \vec{w}^T \vec{g}(\vec{x}_n), \beta^{-1})$

Now, use ML to determine  $\tilde{\omega}_{ML}$  &  $\beta_{ML}$ .

Start with the weights:

$$\frac{\partial}{\partial \omega_j} \log P(\vec{t} | \tilde{\omega}, \beta) = \frac{\beta}{N} \sum_{n=1}^N (t_n - \tilde{\omega}^T \vec{g}(\vec{x}_n)) \times g_j(\vec{x}_n),$$

or

$$\underbrace{\nabla_{\tilde{\omega}} \log P}_{\text{"0}} = \beta \sum_{n=1}^N (t_n - \tilde{\omega}^T \vec{g}(\vec{x}_n)) \vec{g}(\vec{x}_n) = 0$$

$$\Rightarrow \sum_{n=1}^N t_n \vec{g}_j(\vec{x}_n) = \sum_{i=0}^D \omega_i \sum_{n=1}^N g_i(\vec{x}_n) \vec{g}_j(\vec{x}_n)$$

Call  $g_j(\vec{x}_n) = \varphi_{nj}$   $\rightarrow$  elements of  $N \times (D+1)$  design matrix

$$\text{Then } \underbrace{\sum_n t_n \varphi_{nj}}_{\substack{\text{D+1} \\ \text{vector}}} = \sum_i \omega_i \underbrace{\sum_n \varphi_{ni} \varphi_{nj}}_{(D+1) \times (D+1) \text{ matrix}} = \underbrace{\sum_i (\varphi^T \varphi)_{ji} \omega_i}_{\substack{\text{D+1} \\ \text{vector}}}$$

Finally,  $\varphi^T \vec{t} = (\varphi^T \varphi) \tilde{\omega}$ , or

$$\tilde{\omega}_{ML} = (\varphi^T \varphi)^{-1} \varphi^T \vec{t}$$

$\nearrow$        $\equiv$

normal eq's for the least-squares problem

$$\varphi = \begin{pmatrix} g_0(\vec{x}_1) & g_1(\vec{x}_1) & \dots & g_D(\vec{x}_1) \\ \vdots & \vdots & & \vdots \\ g_0(\vec{x}_N) & g_1(\vec{x}_N) & \dots & g_D(\vec{x}_N) \end{pmatrix}$$

$\tilde{\Phi} = (\Phi^T \Phi)^{-1} \Phi^T$  is called the Moore-Penrose pseudo-inverse of  $\Phi$ .

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Next, consider

$$E(\tilde{\omega}) = \frac{1}{2} \sum_{n=1}^N (t_n - \omega_0 - \sum_{j=1}^D w_j \varphi_j(\tilde{x}_n))^2$$

$$\frac{\partial E}{\partial \omega_0} = - \sum_{n=1}^N (t_n - \omega_0 - \sum_{j=1}^D w_j \varphi_j(\tilde{x}_n)) = 0$$

gives

$$\omega_0^{ML} = \frac{1}{N} \sum_{n=1}^N t_n - \frac{1}{N} \sum_{n=1}^N \sum_{j=1}^D w_j \varphi_j(\tilde{x}_n) =$$

$$= \bar{t} - \frac{1}{N} \sum_{n=1}^N \underbrace{\sum_{j=1}^D w_j \varphi_j}_{\frac{1}{N} \sum_{n=1}^N \dots}$$

Thus  $\omega_0^{ML}$  is the difference between the average of the target values and the weighted sum of basis f'n averages.

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Finally, maximize (\*\*) wrt  $\beta$ :

$$\frac{1}{\beta_{ML}} = \underbrace{\frac{1}{N} \sum_{n=1}^N (t_n - \tilde{\omega}_0^{ML} \tilde{\varphi}(\tilde{x}_n))^2}_{\text{residual variance of the target values around } y(\tilde{x}, \tilde{\omega})}$$

residual variance of the target values around  $y(\tilde{x}, \tilde{\omega})$

## Geometry of least squares

$\tilde{t} = (t_1, \dots, t_N)$  is an  $N$ -dim vector

Each  $\tilde{g}_j = (g_j(\tilde{x}_1), \dots, g_j(\tilde{x}_N))$  is also an  $N$ -dim vector

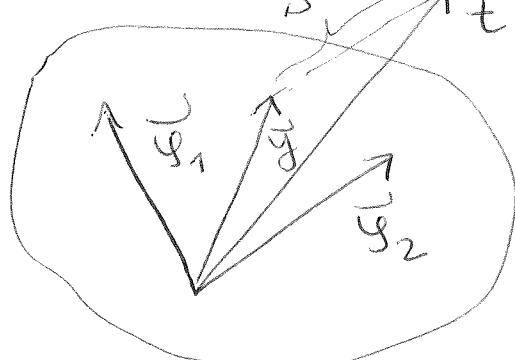
[ $\tilde{g}_j$  is the  $j^{\text{th}}$  column of  $\Phi$ ]

If  $D+1 < N$ , then  $\tilde{g}_j$  vectors

the # of basis functions, including  $g_0(\tilde{x}_n)$

will span a subspace  $S$  of dimensionality  $D+1$  in the  $N$ -dim'l space.

$S$  Euclidean distance between  $\tilde{y}$  &  $\tilde{t}$  Define



$\tilde{y} = (y(\tilde{x}_0, \tilde{w}), \dots, y(\tilde{x}_N, \tilde{w}))$   
N-dim vector, a linear combination of  $D+1$   $\tilde{g}_j$  vectors.

$E(\tilde{w}) \sim \text{second}$  square of Euclidean dist. between  $\tilde{y}$  &  $\tilde{t}$ , so we need to minimize it. This is accomplished by the orthogonal projection of  $\tilde{t}$  into subspace  $S$ .

In practice, opt op may be hard to invert, when some  $\tilde{g}_j$ 's are nearly colinear  $\rightarrow$  add a regularization term to prevent this.

# Regularized least squares

Consider now

$$\tilde{E}(\tilde{\omega}) = \frac{1}{2} \sum_{n=1}^N (t_n - \tilde{\omega}^\top \tilde{\Phi}(\tilde{x}_n))^2 + \frac{\lambda}{2} \tilde{\omega}^\top \tilde{\omega}$$

↑  
still a quadratic fn of  $\omega_j$

$$\tilde{\omega}_{ML} = (\phi^\top \phi)^{-1} \phi^\top \tilde{t} \quad \text{becomes}$$

$$\tilde{\omega}^* = (\lambda \mathbb{I} + \phi^\top \phi)^{-1} \phi^\top \tilde{t}, \quad \begin{matrix} \text{[the inverse matrix} \\ \text{is now} \\ \text{regularized} \end{matrix}$$

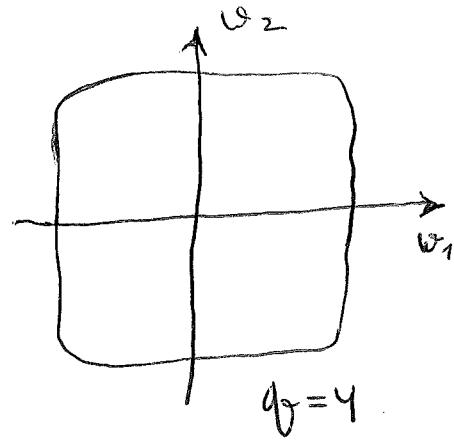
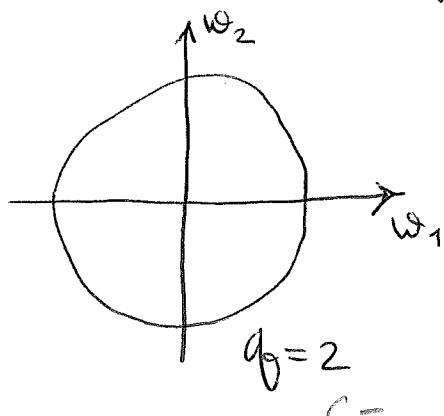
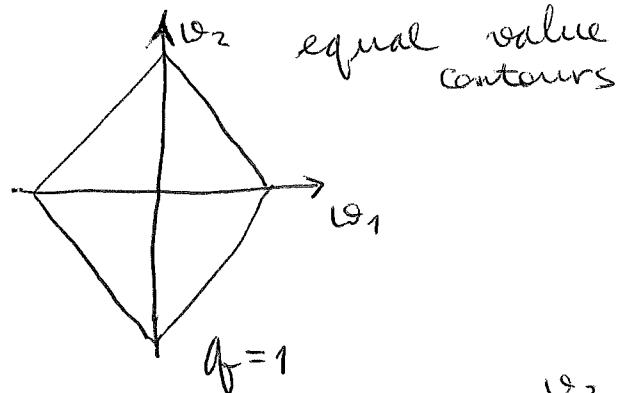
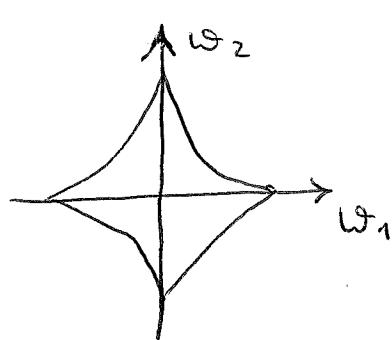
↑ unit matrix  $(D+1) \times (D+1)$

More generally,

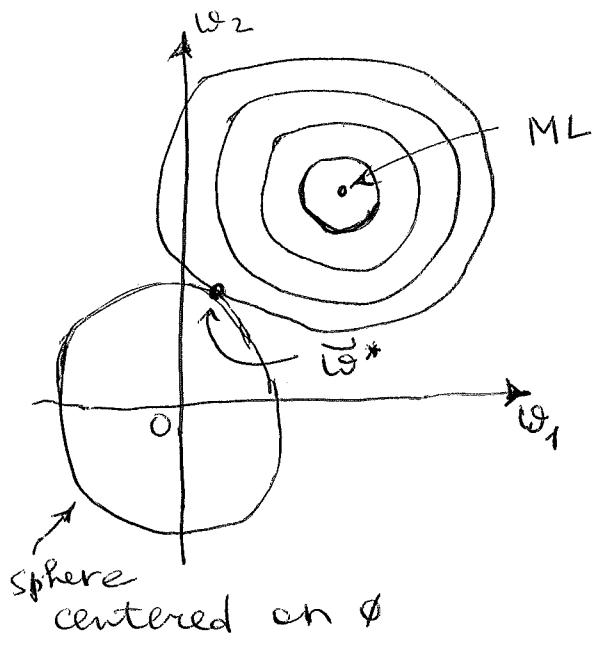
$$\tilde{E}(\tilde{\omega}) = \frac{1}{2} \sum_{n=1}^N (\dots)^2 + \frac{\lambda}{2} \sum_{j=1}^{D+1} |\omega_j|^q$$

$q_f = 2 \leftarrow$  quadratic regularizer above

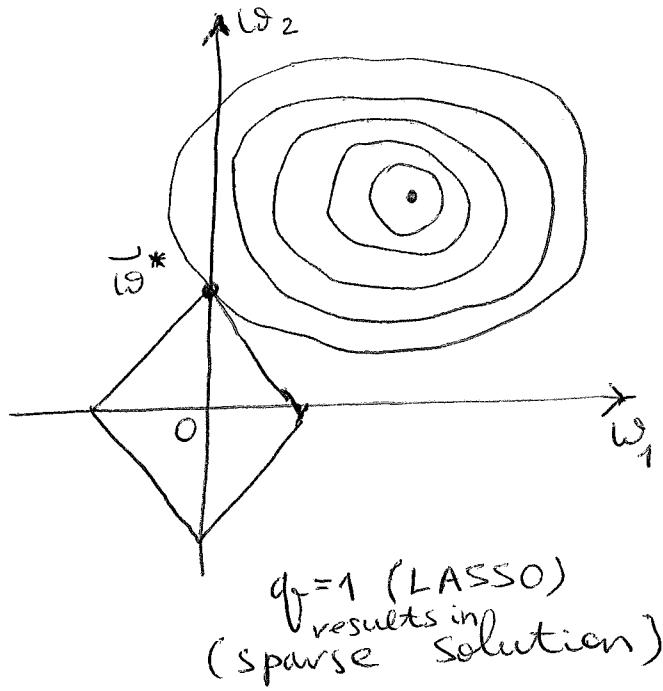
$q_f = 1 \leftarrow$  LASSO regularizer



Now, consider



$$q_f = 2$$



Loss functions for regression (1.5.5)

Consider  $\tilde{x} \rightarrow t$   
input target

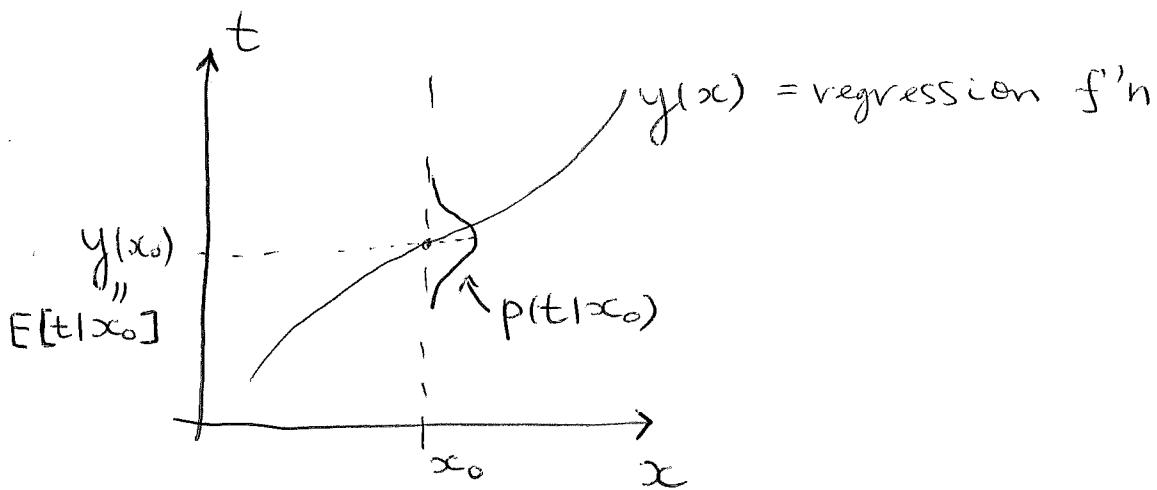
$y(\tilde{x})$  model for  $t$   $\rightarrow L(t, y(\tilde{x}))$  e.g. loss  $= \underbrace{(y(\tilde{x}) - t)^2}_{\text{squared loss}}$

$E[L] = \int d\tilde{x} dt p(\tilde{x}, t) \underbrace{(y(\tilde{x}) - t)^2}_{\text{loss f'n}} \underbrace{p(\tilde{x}, t)}_{\text{joint prob. of } \tilde{x} \& t}$   
expected value of loss

$\frac{\delta E[L]}{\delta y(\tilde{x})} = 2 \int dt (y(\tilde{x}) - t) p(\tilde{x}, t) = 0$   
minimize  $E[L]$

$$y(\tilde{x}) = \frac{\int dt p(\tilde{x}, t) t}{p(\tilde{x})} = \int dt t p(\tilde{x}, t) = E[t | \tilde{x}] \quad (+)$$

$y(\bar{x})$  is called a regression f'n



Further,

$$(y(\bar{x}) - t)^2 = (y(\bar{x}) - E[t|\bar{x}] + E[t|\bar{x}] - t)^2 = \\ = (y(\bar{x}) - E)^2 + 2(y(\bar{x}) - E)(E - t) + (E - t)^2$$

f'n of  $\bar{x}$ , not t

$$\text{Then } E[L] = \int dt d\bar{x} p(\bar{x}, t) \left[ (y(\bar{x}) - \overbrace{E[t|\bar{x}]}^{\text{f'n of } \bar{x}})^2 + \right. \\ \left. + 2(y(\bar{x}) - E[t|\bar{x}])(E[t|\bar{x}] - t) + (E[t|\bar{x}] - t)^2 \right] = \\ = \int d\bar{x} p(\bar{x}) \underbrace{(y(\bar{x}) - E[t|\bar{x}])^2}_{p(t|\bar{x})p(\bar{x})} + \\ + 2 \int d\bar{x} (y(\bar{x}) - E[t|\bar{x}]) \underbrace{\int dt p(\bar{x}, t) (E[t|\bar{x}] - t)}_{E[t|\bar{x}]p(\bar{x}) - p(\bar{x})E[t|\bar{x}]} + \\ \underbrace{2 \text{nd term vanishes}}_{E[t|\bar{x}]p(\bar{x}) - p(\bar{x})E[t|\bar{x}]} = 0 \\ + \int dt d\bar{x} p(\bar{x}, t) (E[t|\bar{x}] - t)^2.$$

Only the 1st term depends on  $y(\bar{x})$ ,  
and will be = 0 (at min) ibb.

$$y(\bar{x}) = E[t|\bar{x}] \text{, consistent with (+)}$$

The 3rd term represents noise in the target values & is indep. of the model:

$$\begin{aligned}
 & \underbrace{\int d\vec{x} p(\vec{x}) E[t|\vec{x}]}_2 - 2 \int dt d\vec{x} t p(\vec{x}, t) E[t, \vec{x}] + \\
 & \cancel{\int dt t^2 p(t)} + \int dt t^2 p(t) = \\
 & = \int d\vec{x} \underbrace{\frac{\int dt t p(\vec{x}, t) \int dt' t' p(\vec{x}, t')}{p(\vec{x})}}_{p(t)} - \langle t^2 \rangle \\
 & \xrightarrow{p(t|\vec{x}) = \frac{p(\vec{x}, t)}{p(\vec{x})}} - 2 \int dt d\vec{x} t p(\vec{x}, t) \underbrace{\frac{\int dt' t' p(\vec{x}, t')}{p(\vec{x})}}_{\langle t \rangle} + \underbrace{\int dt t^2 p(t)}_{\langle t^2 \rangle} = \\
 & = \langle t^2 \rangle - \underbrace{\int \frac{d\vec{x}}{p(\vec{x})} (\int dt t p(\vec{x}, t) \int dt' t' p(\vec{x}, t'))}_{\int d\vec{x} p(\vec{x}) \left( \frac{\int dt t p(\vec{x}, t)}{\int dt p(\vec{x}, t)} \right)^2} = \int d\vec{x} p(\vec{x}) \langle t(\vec{x}) \rangle^2 \\
 & \langle t^2 \rangle = \int dt t^2 p(t) = \frac{\int dt d\vec{x} t^2 p(\vec{x}, t) \langle t \rangle_{\vec{x}}}{\underbrace{\int dt d\vec{x} p(\vec{x}, t)}_{=1}} \Rightarrow \int d\vec{x} \langle t^2(\vec{x}) \rangle p(\vec{x}) \\
 & \langle \dots \rangle \text{ signifies an average wrt } p(\vec{x}, t) \\
 & \langle t^2(\vec{x}) \rangle = \frac{\int dt t^2 p(\vec{x}, t)}{\int dt p(\vec{x}, t)} \\
 & \text{So, } \int dt d\vec{x} p(\vec{x}, t) (E[t|\vec{x}] - t)^2 = \\
 & = \int d\vec{x} p(\vec{x}) [\langle t^2(\vec{x}) \rangle - \langle t(\vec{x}) \rangle^2] = 
 \end{aligned}$$